

Chapter 2

The Monte-Carlo Method and applications to option pricing

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2.1 The Monte-Carlo method

The basic principle of the Monte-Carlo method is to implement on a computer the strong law of large numbers (SLLN): if $(X_k)_{k \geq 1}$ is a sequence, defined on a prob. space $(\Omega, \mathcal{F}, P)$, of independent random variables with the same distribution as X , then:

$$P(dw) - a.s. \quad \bar{X}_n = \frac{X_1(\omega) + \dots + X_n(\omega)}{n} \xrightarrow{n \rightarrow \infty} m_X = E(X)$$

The sequence (X_k) is also called an i.i.d. sequence of random variables with distribution $\mu = P_X$.

Conditions to implement this MC method on a computer:

- * generate pseudo-random numbers $(U_k)_{k \geq 1}$ of law $\mathcal{U}(0,1)$
- * generate X as a function of $U_1, U_2, \dots$:
 $X = \varphi_\tau(U_1, \dots, U_\tau)$ (τ is a finite stopping rule)

2.1.1. Rates of convergence

The (weak) rate of convergence in the SLLN is governed by the Central Limit Theorem (CLT) which says that if $X \in L^2(\mathbb{R})$ then:

$$\sqrt{n} (\bar{X}_n - m_X) \xrightarrow[n \rightarrow \infty]{law} \mathcal{N}(0; \sigma_X^2)$$

where $\sigma_X^2 = \text{Var}(X)$

the mean quadratic error is exactly:

$$\begin{aligned} \|\bar{X}_n - m_x\|_2 &= E((\bar{X}_n - m_x)^2)^{1/2} \\ &= \frac{\sigma_x}{\sqrt{n}} \end{aligned}$$

$$\begin{aligned} \text{indeed: } E((\bar{X}_n - m_x)^2) &= E\left(\left(\frac{\sum_{i=1}^n X_i - m_x}{n}\right)^2\right) \\ &= \frac{\sum_{i=1}^n E((X_i - m_x)^2)}{n^2} \\ &= \frac{n\sigma_x^2}{n^2} = \frac{\sigma_x^2}{n} \end{aligned}$$

This error is also known as the RMSE (Root Mean Square Error)

2.1.2. Confidence level and confidence interval

The CLT also reads: $\Gamma_n\left(\frac{\bar{X}_n - m_x}{\sigma_x}\right) \xrightarrow[n \rightarrow \infty]{\text{loi}} \mathcal{N}(0,1)$

so for all $a, b \in \mathbb{R}$ ($a < b$):

$$\lim_{n \rightarrow \infty} P\left(\Gamma_n\left(\frac{\bar{X}_n - m_x}{\sigma_x}\right) \in [a, b]\right) = \Phi(b) - \Phi(a)$$

where Φ is the cumulative distribution function of $\mathcal{N}(0,1)$

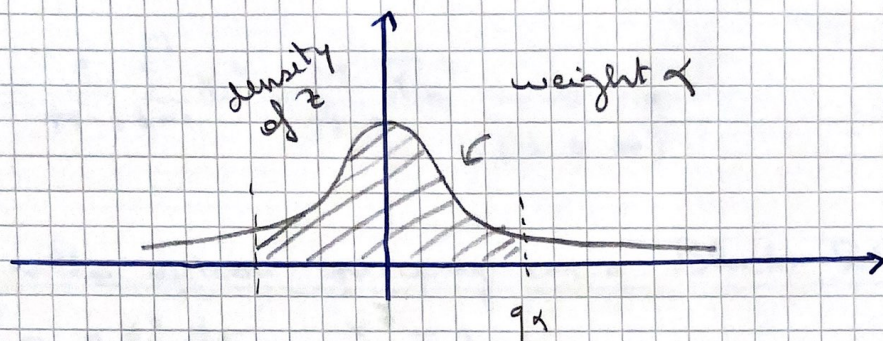
$$\Phi(x) = \int_{-\infty}^x \frac{e^{-t^2/2}}{\sqrt{2\pi}} dt$$

We usually assume $\Gamma_n\left(\frac{\bar{X}_n - m_x}{\sigma_x}\right)$ is of law $\mathcal{N}(0,1)$.

Let $\alpha \in (0,1)$ denote a confidence level (close to 1) and let q_α be the two-sided quantile defined as the unique sol. to:

$$P(|Z| \leq q_\alpha) = \alpha \quad \text{i.e.} \quad 2\Phi(q_\alpha) - 1 = \alpha$$

($Z \sim \mathcal{N}(0,1)$)



Small computation: we want $\mathbb{P}(|Z| \leq q_\alpha) = 1 - \alpha$,

This is: $\mathbb{P}(Z > q_\alpha) + \mathbb{P}(Z < -q_\alpha) = 1 - \alpha$

These two are equal by symmetry

$$2 \mathbb{P}(Z > q_\alpha) = 1 - \alpha$$

$$2(1 - \Phi(q_\alpha)) = 1 - \alpha$$

$$2\Phi(q_\alpha) - 1 = \alpha$$

$$\Phi(q_\alpha) = \frac{1 + \alpha}{2}$$

We take $a = q_\alpha$ and $b = -q_\alpha$ and we can define a theoretical confidence interval at level α :

$$J_M^\alpha = \left[\bar{X}_M - q_\alpha \times \frac{\sigma_x}{\sqrt{M}}, \bar{X}_M + q_\alpha \frac{\sigma_x}{\sqrt{M}} \right]$$

which satisfies: $\mathbb{P}(m_x \in J_M^\alpha)$

$$\mathbb{P}\left(\frac{1}{\sqrt{M}} \left| \bar{X}_M - m_x \right| \leq q_\alpha \right)$$

$$(M \rightarrow \infty) \mathbb{P}(|Z| \leq q_\alpha) = 1 - \alpha$$

However: J_M^α is theoretical because we do not know σ_x .

We usually replace σ_x^2 by its Monte-Carlo estimate:

$$\begin{aligned} \bar{V}_M &= \frac{1}{M-1} \sum_{k=1}^M (X_k - \bar{X}_M)^2 = \frac{1}{M-1} \sum_{k=1}^M X_k^2 + \frac{M}{M-1} \bar{X}_M^2 \\ &\quad - \frac{2\bar{X}_M}{M-1} \sum_{k=1}^M X_k \end{aligned}$$

$$\dots = \frac{1}{n-1} \sum_{i=1}^n X_i^2 - \frac{n}{n-1} \bar{X}_n^2 \xrightarrow{(n \rightarrow \infty)} \sigma_x^2$$

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2.1.3 Vanilla option pricing in a Black-Scholes model

$$\begin{cases} dX_t^0 = rX_t^0 dt, & X_0^0 = 1 \\ dX_t^1 = X_t^1 (r dt + \sigma_1 dW_t^1), & X_0^1 = x_0^1 \\ dX_t^2 = X_t^2 (r dt + \sigma_2 dW_t^2), & X_0^2 = x_0^2 \end{cases}$$

$W = (W^1, W^2)$ is a bi-dim. Brownian motion such that $\langle W^1, W^2 \rangle_t = \rho t$ ($\rho \in [-1, 1]$).

In other words, we can decompose

$W_t^2 = \rho W_t^1 + \sqrt{1-\rho^2} \tilde{W}_t^2$ where $(W^1, \tilde{W}^2)$ is a standard 2-dim. Brownian motion.

filtration: $(F_t)_{t \in [0, T]} = \sigma(W_s, 0 \leq s \leq t, \mathcal{V}_P)$
 augmented with $\mathbb{P}$ -negligible sets

$\mathbb{P}$ is the risk-neutral probability, meaning that $e^{-rt} X_t = e^{-rt} \begin{pmatrix} X_t^1 \\ X_t^2 \end{pmatrix}$ is a $(\mathbb{P}, (F_t))$ -martingale

For all t : $X_t^0 = e^{rt}$, $X_t^i = x_0^i \exp\left[\left(r - \frac{\sigma_i^2}{2}\right)t + \sigma_i W_t^i\right]$ ($i=1, 2$)

(this is an application of Itô's Lemma)

A European vanilla option with maturity $T > 0$ is an option related to a European payoff:

$$h_T = h(X_T)$$

which only depends on X at time T . In such a complete market, the option premium at time $t \in [0, T]$ is given by $V_t = e^{-r(T-t)} \mathbb{E}(h(X_T) | F_t)$
 $(V_0 = e^{-rT} \mathbb{E}(h(X_T)))$

As W has independent stationary increments, we have:

$$\left(\frac{X_T^i}{X_t^i} \right)_{i=1,2} \stackrel{\text{law}}{=} \left(\frac{X_{T-t}^i}{X_0^i} \right)_{i=1,2}$$

We define $v(x_0, T) = e^{-\alpha T} \mathbb{E}(h(X_T))$ and we compute

$$\begin{aligned} V_t &= e^{-\alpha(T-t)} \mathbb{E}(h(X_T) | \mathcal{F}_t) \\ &= e^{-\alpha(T-t)} \mathbb{E}\left(h\left[\left(\frac{X_T^i}{X_t^i}\right)_{i=1,2}\right] | \mathcal{F}_t\right) \\ &= e^{-\alpha(T-t)} \left[\mathbb{E} h\left(\left(x^i \times \frac{X_{T-t}^i}{x_0^i}\right)_{i=1,2}\right) \right]_{x^i = X_t^i, i=1,2} \\ &= v(X_t, T-t) \end{aligned}$$

Examples: * Vanilla call with strike price K :

$$h(x^1, x^2) = (x^1 - K)_+$$

There is a closed form: Black-Scholes formula.
(for V_t)

* Best-of-call with strike price K : $h_T = (\max(X_T^1, X_T^2) - K)_+$

↳ a quasi-closed form exists but we will price it using a Monte-Carlo simulation

$$\text{We re-write: } e^{-\alpha T} h_T \stackrel{\text{law}}{=} Y(z^1, z^2)$$

$$\begin{aligned} &:= \left[\max \left(x_0^1 \exp\left(-\frac{\sigma_1^2}{2}T + \sigma_1 \sqrt{T} z^1\right), \right. \right. \\ &\quad \left. \left. x_0^2 \exp\left(-\frac{\sigma_2^2}{2}T + \sigma_2 \sqrt{T} (\sqrt{\rho} z^1 + \sqrt{1-\rho} z^2)\right) \right) \right. \\ &\quad \left. - K e^{-\alpha T} \right]_+ \end{aligned}$$

where $Z = (z^1, z^2) \sim \mathcal{N}(0, I_2)$

We take a n -sample $(Z_m)_{1 \leq m \leq n}$ and get

$$\begin{aligned} (\text{Best-of-call}) &= \mathbb{E}(Y(z^1, z^2)) \\ &\simeq \bar{Y}_n := \frac{1}{n} \sum_{m=1}^n Y(Z_m) \end{aligned}$$

Estimate for the variance:

$$\bar{V}_M(\varphi) = \frac{1}{M-1} \sum_{m=1}^M \varphi(z_m)^2 - \frac{M}{M-1} \bar{\varphi}_M^2 \approx \text{Var}(\varphi(z)) \quad \left| \frac{14}{17} \right.$$

(for M large enough).

Level of confidence: $\alpha \in (0, 1)$.

Confidence interval:

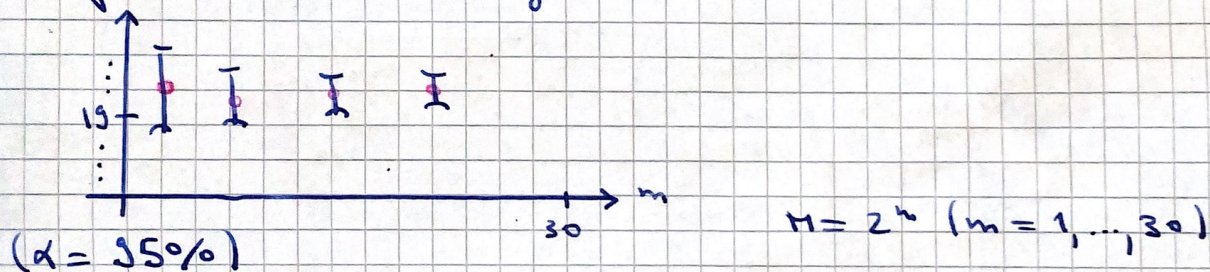
$$I_M^\alpha = \left[\bar{\varphi}_M - q_\alpha \sqrt{\frac{\bar{V}_M(\varphi)}{M}}, \bar{\varphi}_M + q_\alpha \sqrt{\frac{\bar{V}_M(\varphi)}{M}} \right]$$

where q_α is defined by: $2\Phi(q_\alpha) - 1 = \alpha$.

Numerical application:

$\alpha = 0.1$, $r_i = 0.2 = 20\%$, $\bar{r} = 0.5$, $X_0^i = 100$, $T = 1$, $K = 100$

Idea of what one should get



Methodology

Monte-Carlo simulation to compute $m_X = \mathbb{E}(X)$

- * specification of a confidence interval $\alpha \in (0, 1)$
- * simulation of a M sample, computation of empirical mean $\bar{X}_M$ and variance $\bar{V}_M$
- * computation of the resulting confidence interval I_M at confidence level α